

Market Interactions in Gold and Stock Markets: Evidences from Saudi Arabia

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ABSTRACT

The price movements in gold market are considered to detect non-linear dependencies with stock market in the Saudi Arabian context. Both the univariate and multivariate models of generalized autoregressive conditional heteroskedasticity (GARCH) class are employed in this study. Initially, the work uses GARCH (1,1) specification to detect the persistence level of volatility. Proceeding further, a series of models are used to study leverage effect, spillover pattern, risk-premium effects, absolute returns and power transformation factors, etc., Finally, diagonal Baba, Engle, Kraft and Kroner specification is used to determine the contagion effect between gold and stock markets. The findings chiefly prove that a dynamic relationship between gold and stock market do not exist.

Keywords: Gold Return, Multivariate Generalized Autoregressive Conditional Heteroskedasticity, Market Spillover, Contagion Effect, Volatility Persistence JEL Classifications: C58, G11

1. INTRODUCTION

Studying asset price movements continues to be a dynamic area in financial research. This owes to two facts; one is lack of consensus about the short and long term behavior of the asset prices in various markets over different time periods. The other is the advances in powerful econometric tools capable of investigating the market movements. However, unlike pure financial assets such as stocks, bonds or currencies, the gold market received less attention in terms of analyzing the price and return behavior. Well standardized international gold market exhibits data patterns similar to stock or bond markets. Gold being an asset showing attributes of a commodity and a financial asset requires further analysis in the domains of price volatility, persistence and spillover behaviors.

In addition to the attribute of gold as a secure asset, gold price is used as an economic indicator especially in predicting inflation (Garner, 1995; Koeing, 1996; Haubrich, 1998 and Sam and Li, 2004).

From empirical perspectives, there are evidences of long memory in gold returns (GR) (Booth et al. (1982). On the other hand,

Ball et al. (1982) identified the presence of short-term memory. The problem of conditional heteroscedasticity is also displayed in gold data (Akgiray et al., 1991; Frank and Thanasis; 1989). Further, there is a trend among investors to produce a hedge to stay away from risks associated with stock markets by investing in oil and precious metals like gold which have different volatilities and returns of lesser correlation with stocks. Daskalaki and Skiadopoulos (2011) report that due to increasing instability and uncertainty in stock markets investors looks to metal markets as safe avenue. Baur and Lucey (2010) found that gold serves as a safe alternative against stocks in the UK, Germany and US particularly following severe adverse shocks on stock markets. They also found for the period 1979 to 2009 that gold was also a hedge for stocks in UK and US. In a similar study Baur and McDermott (2010) found that during the period 1979 to 2009 gold served as a secured alternative against stocks of major US and European markets but not for Japan, Canada, Australia and BRIC countries.

The issues related to gold as a hedging instrument, price discoverer and safer investment avenue still continue to be a point of interest for researchers as well as practitioners. The evidences are mostly market driven and time specific besides varying on

methodological choices. In this backdrop, it becomes relevant to examine the market interaction between two asset classes of gold and stock in a relatively emerging and under- researched market like Saudi Arabia. The chosen market is the largest in size in the region. Further, considering the nature of data an appropriate set of methodology is employed to investigate the matter under consideration. Initially, generalized autoregressive conditional heteroskedasticity (GARCH) (1,1) specification is used to detect the persistence level of volatility in the markets. GARCH class models are further applied to study leverage effect, spillover pattern, risk-premium effects, absolute returns and power transformation factors etc. Finally, diagonal Baba, Engle, Kraft and Kroner (BEKK) specification is used to determine the contagion effect between gold and stock markets. The findings derived using robust econometric tools on a typical less-explored market are expected to provide new insights into the area and also strengthen the existing level of knowledge. The remainder of the paper is organized as follows. Section 2 presents a brief review of literature related to gold price and return. Methodology and data specification are presented in some detail in Section 3 and Section 4 respectively. Section 5 is an account of results and discussions and finally section 6 offers conclusions followed by references.

2. LITERATURE REVIEW

This section briefly reviews a few important works carried out in the domain of gold as an asset class. The methods adopted are highlighted. Contuk et al. (2013) while studying the effect of fluctuation of gold price on stock prices found an ARCH effect in both gold price and stock price and proposed for GARCH modeling. Their results showed that both the variables were affected by shock of others and also by their own shocks. Arouri et al. (2012) using different multivariate GARCH models found significant return and volatility cross effects between gold prices and stock prices in China. More specifically past GR has an important role in explaining the dynamics of conditional return and volatility of Chinese stock market and hence must thus be taken into consideration while predicting future stock returns. In fact, addition of gold into a portfolio of Chinese stocks improves the risk-adjusted return and exposures related to gold risk can be successfully hedged in portfolios of stocks. Also, the study suggest that amongst, BEKK-CCC/DCC, diagonal BEKK, and VAR-GARCH, the VAR-GARCH model works better than the other. Mishra et al. (2010), using Granger Causality test not only found a causal relationship between stock market returns and gold prices in India but also were important for predicting each other.

But there have been some studies which have different findings. Smith (2001) in a study on US using cointegration and Granger Causality found a slightly inverse relation between the gold price and stock prices. In a related study on Europe and Japan, Smith (2002) using cointegration and Granger Causality found that in the short term there was a negative relation between gold prices and the stock market in the short run, but in the long term the relationship was not significant. While the degree and direction of relationship is still debatable another important dimension from which this relationship should be seen through is from the perspective of relatively abundant oil exporting countries as the country of interest in this study is Saudi Arabia. Commodity prices have a tendency to move together as they are influenced by general macroeconomic factors like interest rates, exchange rates and inflation (Hammoudeh et al., 2008). Gold has an important place amid the main precious metal and is even taken to be the leader of precious metals as there is a parallel movement between gold and other precious metals (Sari et al., 2010). Investors from advanced and emerging markets frequently move between gold and oil and also combine them to diversify their portfolios (Soytas et al., 2009). Arouri and Nguyen (2010) found that oil prices have a tendency to influence various sectors like, oil and gas, financials, industrials and utilities, although but, the degree and the direction of the effect was different for different sectors. In this respect the effect of oil prices for oil exporting countries becomes interesting.

An increase in oil price will have a positive effect in an oil-exporting country, as the income of the country increases which would lead to an increase in expenditure and investments, enhancing productivity and employment. As a result of all of these the stock markets responds positively. Kim and Dilts (2011) are of the opinion that gold and oil serve as substitutes to investments in US dollar value as they are safer. Mazin et al. (2010) found that oil prices do not have a tendency of affecting GCC stock markets and hence cannot be used as predictors. Arouri and Rault (2012) found that oil price shocks of positive nature had a positive impact on the stock market performance of GCC countries. Hammoudeh and Aleisa (2004) and Bashar (2006) found a bidirectional relationship between stock markets and oil prices, in oil-exporting countries. Hammoudeh and Li (2008) opined that major events that lead to changes in oil prices have a propensity to increase the volatility of stock market in GCC countries.

3. METHODOLOGY

A broad category of methods have been employed to examine the issue of price dependencies of various asset classes. Distributional properties of asset returns largely follow stochastic process and are often different from normality (Cheung and Kon, 1993). There are evidences of non-stationarity in the return distributions (Jacky, 1987; Ahn and Reinsel, 1990), Antoniou et al. (2005). Considering the distinctive features of the financial time series such as volatility clustering, leverage effects, persistence and memory levels of volatility, large amount of empirical attempts applied models of ARCH/GARCH family initiated by Engle (1982) and popularized by Bollerslev (1986) and Taylor (1986). However, empirical results show that simple parameterization of a particular type of GARCH models need not necessarily accomplish the objective of describing asset returns (Andersen et al., 2006; Bauwens et al., 2006). Therefore, we apply major classes of GARCH model and examine their predictive abilities. In this empirical work, we use a variety of GARCH models, as they are found to be the appropriate models in studying financial asset prices and related areas. We begin with basic GARCH (p, q) specification in order to detect the volatility pattern in the gold market and stock market. The model specification is given as follows.

$$h_{t} = \alpha_{0} + \sum_{i=1}^{q} \alpha_{i} \varepsilon_{t-i}^{2} + \sum_{j=1}^{p} \beta_{j} h_{t-j}$$
(1)

Where α_0 captures the long-run average volatility. The terms α and β are the ARCH and GARCH coefficients; q is the order of the moving average ARCH terms and p is the number of the autoregressive GARCH terms.

Notably, standard GARCH model does not allow to measure asymmetric behaviour of the volatility parameter. As we are interested in understanding how volatility responds to good and bad news, following Abhyankar (1998), we apply exponential GARCH (EGARCH) specification popularized by Nelson (1991). A general conditional variance framework of EGARCH (1,1) is shown below.

$$\ln\left(h_{t}\right) = \omega + \theta\left[\frac{\varepsilon_{t-1}}{\sqrt{h_{t-1}}}\right] + \gamma\left[\frac{\varepsilon_{t-1}}{\sqrt{h_{t-1}}}\right] - E\left(\frac{\varepsilon_{t-1}}{\sqrt{h_{t-1}}}\right)\right] + \varphi\ln\left(h_{t-1}\right)$$
(2)

According to the EGARCH representation, the variance is conditional on its own past values as well as on past values of the standardized residuals. The persistence of volatility is measured by ϕ for GR. The conditional variances are finite if $\phi < 1$. The coefficients of γ measures the ARCH effects and θ capture the asymmetric behavior. The coefficients of ARCH term, γ determines the spillover behavior of the markets in such a way that a significant negative value indicates that for a given decrease in volatility of the other market. Leverage term is expected to be negative to convey that market reacts more to negative news than to positive (good) news. The lag truncation length (p and q) is determined using likelihood ratio (LR) tests and we choose EGARCH (1,1) specification.

GARCH in Mean and Power GARCH are also used to study the risk premium effects and impact of absolute returns and power transformation. The model frameworks are provided below for both GARCH M and PARCH specifications respectively in equation (3) and equation (4).

$$\sigma_{t}^{2} = \omega + \sum_{j=1}^{q} \beta_{j} \sigma_{t-j}^{2} + \sum_{i=1}^{p} \alpha_{i} \varepsilon_{t-i}^{2} + Z_{t}^{'} \pi$$
(3)

$$\sigma_t^{\delta} = \omega + \sum_{j=1}^q \beta_j \sigma_{t-j}^{\delta} + \sum_{i=1}^p \alpha_i (|\epsilon_{t-i}| - \gamma_i \epsilon_{t-i})^{\delta}$$

$$\tag{4}$$

In the multivariate GARCH class of models, many developments have been reported since the first VEC model of Bollerslev et al. (1988). The VEC model presumes that every conditional variance and covariance is a function of all lagged conditional variances and covariances, as well as lagged squared returns and cross-products of returns. The model is shown below:

$$vech(H_t) = C + \sum_{j=1}^{q} A_j vech(\varepsilon_{t-j}\varepsilon_{t-j}') + \sum_{j=1}^{p} B_j vech(H_{t-j})$$
(5)

Where vech (H_i) is an operator that stacks the columns of the lower triangular part of argument matrix, H_i is the covariance

matrix of the residuals, C is an $N(N+1)/2 \times I$ vector, $A_{j'}$ and B_{j} are $N(N+1)/2 \times N(N+1)/2 \times parameter matrices and <math>\varepsilon$ is an NXI vector.

Baba et al. (1990) introduced a model called BEKK model that ensures positive definiteness. This model can be considered as a restricted version of the VEC model and in this model, a new parameterization of the conditional variance matrix H_t was defined. The BEKK model is given by equation (6).

$$H_{t} = CC' + \sum_{j=1}^{q} \sum_{k=1}^{k} A'_{kj} \varepsilon_{t-j} \varepsilon_{t-j}' A_{kj} + \sum_{j=1}^{p} \sum_{k=1}^{k} B'_{kj} H_{t-j} B_{kj}$$
(6)

Where H_t represents the conditional variance-covariance matrices; A_{kj} , B_{kj} and C are NXN parameter matrices and C' is a lower triangular matrix.

The first order BEKK model is given by equation (7).

$$H_t = CC' + A'\varepsilon_{t-1}\varepsilon'_{t-1}A + B'H_{t-1}B$$
(7)

The BEKK model has also its diagonal form by assuming A_{kj} , and B_{kj} are diagonal matrices. The BEKK model is not linear in parameters. In this model, the elements of the symmetric matrix A measures the degree of market shocks from stock market to gold market and the elements of the symmetric matrix B indicate the persistence in conditional volatility between the two markets. The model ensures that the conditional variance-covariance matrices H_i is positive definite if atleast one of C or B is of full rank.

4. DATA SPECIFICATION

Our study is based on gold price and stock index value. The daily sample spans from November 2007 onwards, inclusive for a total of 1604 observations for both series of the Tadawul All Shares Index (TASI) and gold price. TASI prices is taken as a representative sample of stock prices in the middle east region, as the Tadawul, stock exchange of Saudi Arabia is the largest and most active exchange in the Middle East. TASI daily closing price data is acquired from Tadawul¹. The daily gold price adjusted in Saudi Riyal is obtained from Bullion Rates². The selection of both price series is from same market region so as to ensure the logic for comparison. Price figures are converted into daily returns for both TASI and gold samples, calculated as the first difference of the log value of the closing price. Daily closing price series are used for the study. Data synchronization is made in accordance with joint appearance of both data series with a view to eliminate any problem arising out of missing data. Statistical properties of the sample data are given in Table 1.

Gold reports a better daily return than the stock during the period with a value of 0.0289 as against 0.0028. Standard deviation, a light measure of risk is less with the GR compared to stock return. All the four series show the evidence of skewness and the effect of skewness may be negative which describes the departure from the symmetry. The values of kurtosis show more than three in both return series

¹ http://www.tadawul.com.sa/wps.

² http://www.bullion-rates.com/gold/SAR/.

Table 1: Descriptive statistics

| Item | Mean | Median | Maximum | Minimum | Standard deviation | Skewness | Kurtosis | Jarque-Bera | Р |
|-------------|----------|----------|----------|----------|--------------------|----------|----------|-------------|---|
| TASI return | 0.0028 | 0.0758 | 9.0873 | -10.3285 | 1.4577 | -0.9962 | 14.4807 | 9068.726 | 0 |
| Gold return | 0.0289 | 0.0424 | 10.68935 | -9.12407 | 1.358 | -0.3134 | 9.1954 | 2589.958 | 0 |
| TASI | 7058.026 | 6712.970 | 11697.01 | 4130.010 | 1375.906 | -0.7795 | 3.4232 | 174.305 | 0 |
| Gold | 4837.071 | 4939.74 | 7131.62 | 2672.26 | 1208.007 | -0.04367 | 1.6273 | 126.359 | 0 |

TASI: Tadawul all shares index, SD: Standard deviation, GR: Gold return

indicating non-normality behavior. In addition, the Jarque-Bera values reject the null hypothesis that the price and return series are normally distributed as the associated probability values are all zero.

In order to study the properties of data, each time series we use in this study is subjected to a check for stationarity. The Dickey-Fuller test was separately employed for price and return series of both stock index and gold symbols. Table 2 reports the result of unit root test. It is evident that stock and gold price series are difference-stationary.

The unit root hypothesis (i.e., series is non-stationary) is found true in the closing price series, while the return series are stationary in all the cases. Phillips-Perron test for unit root is carried out on original price series and log return series of both stock index and gold market. The test result is summarized in Table 3.

Another test called DF-GLS is also performed as a modified Dickey–Fuller t-test for a unit root in which the series has been transformed by a generalized least-squares regression. The maximum lag length is 24 chosen by Schwarz criterion. The result confirms the finding reported earlier at 1%, 5% and 10% intervals. The figures are omitted for brevity.

In addition to the tests carried out above, we perform Bartlett's periodogram-based test for white noise. Bartlett's test is a test of the null hypothesis that the data come from a white-noise process of uncorrelated random variables having a constant mean and a constant variance (Bartlett, 1955). The result is presented graphically for each series under consideration.

We can see in the first two graphs (Figure 1) for gold and index price series, that the values denoted by dark colored points do appear outside the confidence bands. We also have confirmation from the test statistic, at a p-value of 0.000, that the process is significantly different from white noise. However, last two graphs for return series for gold and index show that the values fairly appear in the confidence bands. The test statistic has p-values of 0.6149 and 0.0003, so we conclude that the process is not different from white noise.

On detecting the non-stationary patterns in the original price series of gold and stock index, different heteroscedastic models from ARCH/GARCH family class are tried in this work. Each model checks out the interdependencies between the gold market and stock market during the period.

5. RESULTS AND DISCUSSIONS

This section deals with various results and findings obtained out of the estimation process. The results are divided among the areas

Table 2: Unit root test

| Item | Test statistic | Р |
|--------------------------|----------------|--------|
| Gold Z (t) | -1.590 | 0.4885 |
| Stock index Z (t) | -1.304 | 0.6272 |
| Gold return Z (t) | -39.722 | 0.0000 |
| Stock index return Z (t) | -36.656 | 0.0000 |

*Critical values at 1%=-3.430, 5%=-2.860 and 10%=-2.570, GR: Gold return

Table 3: Phillips-Perron test for unit root

| Item | Test | 1% | 5% | 10% |
|-----------------|----------------|--------------------|----------|----------|
| | statistic | critical | critical | critical |
| | | value | value | value |
| Gold | | | | |
| Z (rho) | -3.133 | -20.700 | -14.100 | -11.300 |
| Z (t) | -1.571 | -3.430 | -2.860 | -2.570 |
| MacKinnon appro | oximate p-valu | the for $Z(t)=0.4$ | 4983 | |
| TASI | | | | |
| Z (rho) | -5.201 | -20.700 | -14.100 | -11.300 |
| Z (t) | -1.541 | -3.430 | -2.860 | -2.570 |
| MacKinnon appro | oximate p-valu | the for $Z(t)=0$. | 5134 | |
| Gold return | | | | |
| Z (rho) | -1530.439 | -20.700 | -14.100 | -11.300 |
| Z (t) | -39.741 | -3.430 | -2.860 | -2.570 |
| MacKinnon appro | oximate p-valu | the for $Z(t)=0$. | 0000 | |
| TASI return | | | | |
| Z (rho) | -1480.047 | -20.700 | -14.100 | -11.300 |
| Z (t) | -36.694 | -3.430 | -2.860 | -2.570 |

MacKinnon approximate p-value for Z (t)=0.0000, TASI: Tadawul all shares index, GR: Gold return

of volatility clustering, persistence, asymmetric behavior and price interdependence and contagion effects.

5.1. Volatility Modeling

On detecting the ARCH effects, we employed a series of models in order to investigate the volatility behavior of the market and also to determine the effectiveness of each model. For the GARCH (p, q) model, the values of p and q are set one according to the Schwarz Information Criterion. The result of GARCH (1,1) estimation for index return and GR is presented in Table 4.

It is clear that the coefficients of GARCH constant (α_0), ARCH (α_1) and GARCH (β_1) are significantly different from zero and are within the parametric restrictions. This implies greater impact of shocks (news) on market returns and presence of volatility clustering. A large shock in a day leads to a large (conditional) variances the following day, indicating the impact of recent news on price changes. Again, the higher value of GARCH implies long memory of the model to the shocks. More importantly, the GR is found unaffected by index returns giving a notion that both move independently. The findings also discard the notion of market efficiency because in an efficient market, the gold prices follow stochastic movements leaving no scope for historical analysis

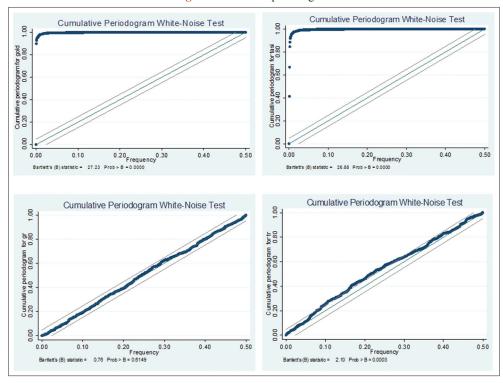


Figure 1: Bartlett's periodogram

Table 4: GARCH estimation output

| | Dependent v | ariable: GR (Gold return) | | |
|------------------------------|-----------------|---------------------------------|-------------|--------|
| | Method: ML-ARCH | (Marquardt)-Normal distribution | | |
| | Convergence | achieved after 36 iterations | | |
| | Presample varia | nce: Backcast (Parameter=0.7) | | |
| Variable | Coefficient | Standard error | z-Statistic | Р |
| С | 0.088882 | 0.022375 | 3.972417 | 0.0001 |
| TR | 0.021905 | 0.017389 | 1.259706 | 0.2078 |
| Variance equation | | | | |
| С | 0.030227 | 0.002675 | 11.30025 | 0.0000 |
| $RESID(-1)^2$ | 0.146044 | 0.013614 | 10.72786 | 0.0000 |
| GARCH(-1) | 0.851042 | 0.010872 | 78.27564 | 0.0000 |
| TR | 0.006277 | 0.009784 | 0.641545 | 0.5212 |
| \mathbb{R}^2 | -0.001755 | Mean dependent var | 0.002814 | |
| Adjusted R ² | -0.002381 | S.D. dependent var | 1.457762 | |
| Standard error of regression | 1.459497 | Akaike info criterion | 2.963493 | |
| Sum squared resid | 3410.339 | Schwarz criterion | 2.983629 | |
| Log likelihood | -2369.240 | Hannan-Quinn criter | 2.970969 | |
| Durbin-Watson stat | 1.820959 | | | |

GR: Gold return, GARCH: Generalized autoregressive conditional heteroskedasticity

to reap any abnormal profit. Asset returns do not exhibit perfect random walk behavior. Many of the chosen factors including own past information set explain them. Figure 2 depicts the volatility patters of gold market during the sample period. The model does a good job in capturing volatility clustering that has occurred at different intervals as is seen in the plot.

5.2. Asymmetric Behavior of Volatility

EGARCH specification provides a structure of asymmetric nature of market volatility. The appropriate model appears to be EGARCH (1,1) and the estimation output is given in the Table 5. In the model specification, all parameters are significant except the index return coefficient. The persistence of volatility is visible

and also the asymmetric nature. The gold market volatility is a function of past innovations. A negative value of asymmetric parameter suggests that the market is rather sensitive to bad information than to positive ones. This is a general trend in asset markets.

5.3. Risk Premium Effects

The return of a security may depend on its volatility (risk). This kind of risk-return relationship is established through GARCH- in- mean (GARCH- M) model which adds a heteroscedasticity term into the mean equation. The mean equation estimation outputs are separately shown for standard deviation and variance in the Tables 6 and 7. Variance equation parameters

Table 5: EGARCH estimation output

| | Dependent variable: GR (Gold return) | | | | | | |
|------------------------------|---|-------------------------------|-------------|--------|--|--|--|
| | Method: ML-ARCH (Marquardt)-Normal distribution | | | | | | |
| | Convergence | achieved after 34 iterations | | | | | |
| | Presample varia | nce: Backcast (Parameter=0.7) | | | | | |
| Variable | Coefficient | Standard error | z-Statistic | Р | | | |
| С | 0.076501 | 0.020665 | 3.701969 | 0.0002 | | | |
| GR | 0.013101 | 0.014361 | 0.912250 | 0.3616 | | | |
| Variance equation | | | | | | | |
| C(3) | -0.215157 | 0.014745 | -14.59156 | 0.0000 | | | |
| C(4) | 0.302894 | 0.019546 | 15.49610 | 0.0000 | | | |
| C(5) | -0.132599 | 0.012312 | -10.76948 | 0.0000 | | | |
| C(6) | 0.960762 | 0.003359 | 286.0041 | 0.0000 | | | |
| C(7) R ² | 0.012856 | 0.011854 | 1.084581 | 0.2781 | | | |
| \mathbb{R}^2 | -0.001416 | Mean dependent var | 0.002814 | | | | |
| Adjusted R ² | -0.002041 | S.D. dependent var | 1.457762 | | | | |
| Standard error of regression | 1.459249 | Akaike info criterion | 2.918158 | | | | |
| Sum squared resid | 3409.183 | Schwarz criterion | 2.941649 | | | | |
| Log likelihood | -2331.903 | Hannan-Quinn criter | 2.926880 | | | | |
| Durbin-Watson stat | 1.822071 | | | | | | |

GR: Gold return

Table 6: GARCH M-SD

| Variable | Coefficient | Standard | z-Statistic | Р |
|-----------------------|-------------|----------|-------------|--------|
| | | error | | |
| (ARCH) ^{1/2} | 0.013686 | 0.066381 | 0.206178 | 0.8367 |
| C | 0.077578 | 0.056387 | 1.375812 | 0.1689 |
| GR | 0.021829 | 0.017395 | 1.254958 | 0.2095 |

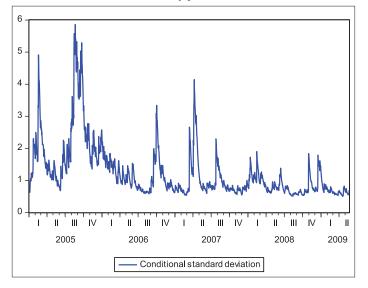
GARCH: Generalized autoregressive conditional heteroskedasticity, SD: Standard deviation

Table 7: GARCH M-variance

| Variable | Coefficient | Standard | z-Statistic | Р |
|----------|-------------|----------|-------------|--------|
| | | error | | |
| GARCH | 0.004568 | 0.020763 | 0.220028 | 0.8258 |
| С | 0.085235 | 0.026872 | 3.171897 | 0.0015 |
| GR | 0.021885 | 0.017384 | 1.258865 | 0.2081 |

GARCH: Generalized autoregressive conditional heteroskedasticity

Figure 2: Generalized autoregressive conditional heteroskedasticity volatility patterns



are not provided, as the results validate the prior findings and also less relevant in the present context.

Symmetric GARCH (1,1)-in -mean estimates show insignificant but positive risk-return relationship in the gold market. The volatility coefficient (risk premium) in both standard deviation (0.0136) and variance (0.0045) neglects a dynamic risk-return pattern. This is against the general notion of a significant risk-return relationship. However, there are evidences for insignificant but positive relationships (Baillie and DeGennarro; 1990).

5.4. Absolute Returns and Power Transformations

A common finding in much of the empirical finance research is that although asset returns contain certain amount of serial correlation, there exists correlation between absolute returns and their power transformations. Ding et al. (1993) investigated the autocorrelation structure of power of the absolute return of daily S and P 500 and found that absolute return has significant positive autocorrelations for long lags. In order to carry out such an analysis on the gold market, we model the mean equation using an MA (1) process. The test result is provided in Table 8.

It is noted that the estimated power GARCH model for the GR exhibits the high level of persistence (0.842). It is concluded that the estimated autocorrelations of the power-transformed absolute observations start high, and decrease very slowly.

5.5. The Contagion Effect between Gold and Stock Markets

According to Forbes and Rigobon (2002), contagion effect is a significant increase in correlation between markets. Such effects are identified in many empirical researches carried out mainly in the area of stock market and crude oil prices. (Filis, 2010; Mollick and Assefa, 2013 and many others). As mentioned earlier, the gold market received less attention in this regard. The link between gold prices and stock market has important implication for investment and hedging.

We make use of diagonal BEKK - GARCH (1,1) and VECH models to study the volatility spillover between gold market and

Table 8: Power ARCH estimates

| Variable | Coefficient | Standard error | z-Statistic | Р |
|------------------------------|-------------|-----------------------|-------------|--------|
| С | 0.078600 | 0.013507 | 5.819394 | 0.0000 |
| GR | 0.004824 | 0.009651 | 0.499813 | 0.6172 |
| Variance equation | | | | |
| C(3) | 0.038053 | 0.003737 | 10.18255 | 0.0000 |
| C(4) | 0.175247 | 0.013437 | 13.04240 | 0.0000 |
| C(5) | 0.585338 | 0.042092 | 13.90616 | 0.0000 |
| C(6) | 0.842900 | 0.009847 | 85.60318 | 0.0000 |
| C(7) | 0.007623 | 0.004769 | 1.598236 | 0.1100 |
| C(8) R ² | 0.648987 | 0.079722 | 8.140624 | 0.0000 |
| \mathbb{R}^2 | -0.002250 | Mean dependent var | 0.002814 | |
| Adjusted R ² | -0.002876 | S.D. dependent var | 1.457762 | |
| Standard error of regression | 1.459857 | Akaike info criterion | 2.896093 | |
| Sum squared resid | 3412.023 | Schwarz criterion | 2.922941 | |
| Log likelihood | -2313.219 | Hannan-Quinn criter. | 2.906062 | |
| Durbin-Watson stat | 1.821262 | | | |

stock market. Both the models are estimated using maximum likelihood techniques. Table 9 reports the BEKK - variancecovariance estimated parameters. The conditional variance of the diagonal BEKK-GARCH model includes two variables namely GR and stock index return. In Table 10, the estimation output of VECH-GARCH specification is presented.

The result shows absence of significant shock transmissions between stock market and gold market. Diagonal BEKK model and VECH model offer similar results that again confirm the findings obtained in the uni-variate GARCH computations. However, there is evidence of persistence of conditional volatility between the two markets. BEKK standard residuals are plotted in Figure 3.

Table 11 shows the test result of GARCH effects on the standardized residuals of the BEKK model. It can be concluded that the GARCH effect has largely disappeared. From the Ljung-Box statistic based on the auto correlation test we tend to accept the null hypothesis that series is the random.

The impulse response graphs plotted in Figure 4 represent how stock index return (TR) responds to shocks of both index equation and gold equation. Similarly, response behavior of gold to shocks of both gold equation and index equation is also depicted in the figure. The line graph in the upper left-hand corner shows that following a shock to the index equation, index twist around for some time, but later the response has effectively dissipated. GR also exhibits similar response to shocks of GR. In contrast, TR and GR are effectively unresponsive to the shocks of each other. This finding reinforces the results obtained in GARCH family estimations. The dynamic and significant relationship among gold price and stock price is absent.

6. CONCLUSIONS

In this paper we have investigated the relationship between gold and stock market. There exists no specific interdependence between the stock price and gold price during the investigated period. Lack of spillover may be because of the fact that any innovation originated in one market is offset by other innovation of opposite sign, leading to a weak or zero effect of stock market on gold market as noted by Jorion (1990) for exchange rate

Table 9: Covariance specification: Diagonal BEKK

| GARCH=M+A1*RESID(-1)*RESID(-1)'*A1+B1*GARCH(-1) |
|---|
| *B1+E1*TR |

M is an indefinite matrix

A1 is a diagonal matrix

B1 is a diagonal matrix

E1 is an indefinite matrix*

| | Transformed Variance Coefficients | | | |
|---------|-----------------------------------|----------|-------------|--------|
| | Coefficient | Standard | z-Statistic | Р |
| | | error | | |
| M(1,1) | 0.033954 | 0.002955 | 11.49083 | 0.0000 |
| A1(1,1) | 0.379818 | 0.013705 | 1.01400 | 0.0085 |
| B1(1,1) | 0.918349 | 0.004714 | 194.8289 | 0.0000 |
| E1(1,1) | -0.085160 | 0.003600 | -0.65802 | 0.0056 |

*Coefficient matrix is not positive semi definite, GARCH: Generalized autoregressive conditional heteroskedasticity, BEKK: Baba, Engle, Kraft and Kroner

Table 10: Covariance specification: Diagonal VECH GARCH=M+A1*RESID(-1)*RESID(-1)'+B1*GARCH(-1)+ E1*TR

- M is a scalar
- A1 is a rank one matrix
- B1 is a rank one matrix

E1 is a rank one matrix

| | Transformed variance coefficients | | | |
|---------|-----------------------------------|----------|-------------|--------|
| | Coefficient | Standard | z-Statistic | Р |
| | | error | | |
| М | 0.231196 | 0.381977 | 0.605261 | 0.5450 |
| A1(1,1) | 0.759816 | 1.249156 | 0.608263 | 0.5430 |
| B1(1,1) | 0.843734 | 0.026030 | 32.41422 | 0.0000 |
| E1(1,1) | 3.14E-05 | 0.138444 | 0.000227 | 0.9998 |

GSRCH: Generalized autoregressive conditional heteroskedasticity

Table 11: Residual portmanteau tests for autocorrelations

| Lags | Q-Stat | Prob. | df |
|------|---------|--------|----|
| 1 | 0.88162 | 0.0006 | 1 |
| 2 | 0.71680 | 0.0004 | 2 |
| 3 | 0.79141 | 0.0013 | 3 |
| 4 | 0.44785 | 0.0016 | 4 |
| 5 | 0.50997 | 0.0024 | 5 |
| 6 | 0.86376 | 0.0044 | 6 |
| 7 | 0.51723 | 0.0046 | 7 |
| 8 | 0.75244 | 0.0054 | 8 |
| 9 | 0.76894 | 0.0096 | 9 |
| 10 | 0.13589 | 0.0036 | 10 |
| 11 | 0.59078 | 0.0004 | 11 |
| 12 | 0.60517 | 0.0008 | 12 |

Figure 3: Baba, Engle, Kraft and Kroner standard residuals

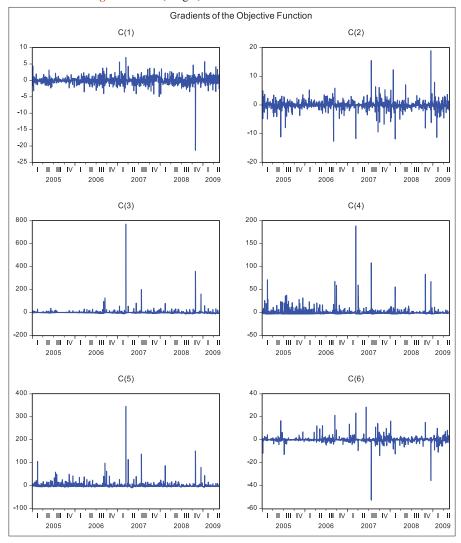
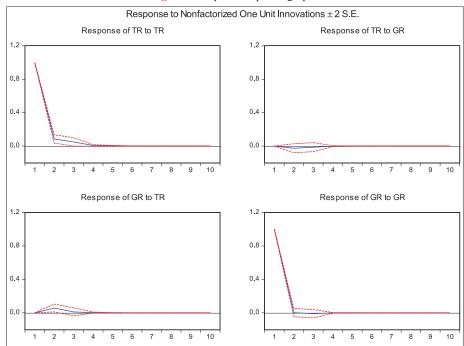


Figure 4: Impulse response graphs



volatility. Another possible explanation is that volatility spillovers are neutralized by the effective use of gold for hedging purpose. Also, a minor chance is that different sets of investors operate in two markets with sole focus on either stock or gold investment. A scrutiny of sample of investor profile may also be required to support any argument.

Risk-premium effect is also absent in the case of gold market. However, an insignificant but positive risk-return relationship in the market can't undoubtedly reject the general hypothesis of risk-return togetherness. Persistence of volatility is found in the market in such a way that market is more responsive to negative effects. A common behavior of asset market wherein low returns are followed by lower returns is found in the case of gold market as well.

In brief, neither a bidirectional price spillover nor a volatility transmission occurs between the gold market and stock market during the study period. Portfolio selection is to be more inclusive to gain the potential benefits of hedging in either market.

REFERENCES

- Abhyankar, A. (1998), Linear and nonlinear granger causality: Evidence from the UK stock index futures market. Journal of Futures Markets, 18(5), 519-540.
- Ahn, S.K., Reinsel, G.C. (1990), Estimation for partially non-stationary multivariate autoregressive models. Journal of the American Statistical Association, 85, 813-823.
- Akgiray, V.G., Booth, G.H., John, J., Chowdhury, M. (1991), Conditional dependence in precious metal prices. The Financial Review, 26(3), 367-386.
- Andersen, T., Bollerslev, T., Christoffersen, P., Diebold, F. (2006), Volatility and correlations forecasting. In: Elliott, G., Granger, C., Timmermann, A., editors. Handbook of Economic Forecasting. Vol. 1. Amsterdam, North-Holland: Elsevier. p778-878.
- Antoniou, A., Koutmos, G., Pericli, A., (2005), Index futures and positive feedback trading: Evidence from major stock exchanges. Journal of Empirical Finance, 12, 219-238.
- Arouri, M., Lahiani, A., Nguyen, D. (2012), World Gold Prices and Stock Returns in China: Insights for Hedging and Diversification Strategies. European Economics and Finance Society Eleventh Annual Conference Koç University, Istanbul.
- Arouri, M., Nguyen, D.K. (2010), Oil prices, stock markets and portfolio investment: Evidence from sector analysis in Europe over the last decade. Energy Policy, 38, 4528-4539.
- Arouri, M., Rault, C. (2012), Oil prices and stock markets in GCC countries: Empirical evidence from panel analysis. International Journal of Finance and Economics, 17(3), 242-253.
- Baillie, R.T., DeGennarro, R.P. (1990), Stock returns and volatility. Journal of Financial and Quantitative Analysis, 25(5), 203-214.
- Ball, C., Torous, W., Tschogel, A. (1982), Gold and the weekend effect. Journal of Futures Markets, 2(2), 175-182.
- Bartlett, M.S. (1955), An Introduction to Stochastic Processes with Special Reference to Methods and Applications. Cambridge: Cambridge University Press.
- Bashar, Z. (2006), Wild oil prices, but brave stock markets! The case of Gulf Cooperation Council (GCC) stock markets. Operational Research: An International Journal, 6(2), 145-162.

Baur, D.G., Lucey, B.M. (2010), Is gold a hedge or a safe haven? An

analysis of stocks, bonds and gold. The Financial Review, 45, 217-229.

- Baur, D.G., McDermott, T.K. (2010), Is gold a safe haven? International evidence. Journal of Banking and Finance, 34, 1886-1898.
- Bauwens, L., Laurent, S., Rombouts, J. (2006), Multivariate GARCH models: A survey. Journal of Applied Econometrics, 21, 79-110.
- Baba, Y., Engle, R.F., Kraft, D., Kroner, K. (1990), Multivariate Simultaneous Generalized Arch, Unpublished Manuscript, University of California, San Diego.
- Bollerslev, T. (1986), Generalized autoregressive conditional heteroscedasticity. Journal of Econometrics, 31, 307-327.
- Bollerslev, T., Engle, R.F., Wooldridge, J.M. (1988), A capital asset pricing model with time varying covariances. Journal of Political Economy, 96, 116-131.
- Booth, G., Kaen, F., Koveos, P. (1982), Persistent dependence in gold prices. Journal of Financial Research, 5(1), 85-93.
- Cheung, Y.W., Kon, S.L. (1993), Do gold market returns have long memory? Financial Review, 28, 181-202.
- Contuk, Y.F., Burucu, H., Gungur, B. (2013), Effect of gold price volatility on stock returns: Example of Turkey. International Journal of Economics and Finance Studies, 5(1), 119-140.
- Daskalaki, C., Skiadopoulos, G.S. (2011), Should investors include commodities in their portfolios after all? New evidence. Journal of Banking and Finance, 35(10), 2606-2626.
- Ding, Z., Granger, C.W.J., Engle, R.F. (1993), A long memory property of stock market returns and a new model. Journal of Empirical Finance, 1, 83-106.
- Engle, R.F. (1982), Autoregressive conditional heteroskedasticity with estimates of the variance of the United Kingdom inflation. Econometrica, 50, 987-1007.
- Filis, G. (2010), Macro economy, stock market and oil prices: Do meaningful relationships exist among their cyclical fluctuations?. Energy Economics, 32(4), 877-886.
- Forbes, K.J., Rigobon, R. (2002), No contagion, only interdependence: Measuring stock and the 2008-2009 financial crisis. Energy Economics, 36, 1-18.
- Frank, M., Thanasis, S. (1989), Measuring the strangeness of gold and silver rates of return. The Review of Economic Studies, 56(4), 553-567.
- Garner, C.A. (1995), How useful are leading indicators of inflation? Economic Review (Federal Reserve Bank of Kansas City), 80(2), 2-7.
- Hammoudeh, S., Li, h. (2008), Sudden changes in volatility in emerging markets: the case of gulf arab stock markets. International review of financial analysis, 17, 47-63.
- Hammoudeh, S., Aleis, E. (2004), Dynamic relationship among GCC stock markets and NYMEX oil futures. Contemporary Economic Policy, 22, 250-269.
- Hammoudeh, S., Sari, R., Ewing, B.T. (2008), Relationships among strategic commodities and with financial variables: A new look. Contemporary Economic Policy, 27(2), 251-264.
- Haubrich, J.G. (1998), Gold prices. Economic commentary. Federal Reserve Bank of Cleveland, 1, 1-4.
- Jacky, C.S. (1987), The sub-gaussian distribution of currency futures: Stable paretian or nonstationary? The Review of Economics and Statistics, 69(1), 100-107.
- Jorion, P. (1990), The exchange rate exposure of US multinationals. Journal of Business, 63, 331-345.
- Kim, M.H., Dilts, D.A. (2011), The relationship of the value of the dollar, and the prices of gold and oil: A tale of asset risk. Economics Bulletin, 31(2), 1151-1162.
- Koeing, E.F. (1996), Should high prices of gold be of concern? The Southwest Economy, Federal Reserve Bank of Dallas, 4, 6-9.

Mazin, A.M.A., Abdulnasser, H.J., Manuchehr, I. (2010), An empirical

investigation of the informational efficiency of the GCC equity markets: Evidence from bootstrap simulation. International Review of Financial Analysis, 19(1), 47-54.

- Mishra, P.K., Das, J.R., Mishra, S.K. (2010), Gold price volatility and stock market returns in India. American Journal of Scientific Research, 9, 47-55.
- Mollick, A.V., Assefa, T.A. (2013), US stock returns and oil prices: The tale from daily data market co-movement. Journal of Finance, 57(5), 2223-2261.
- Nelson, D.B. (1991), Conditional heteroscedasticity in asset returns: A new approach. Econometrica, 59(2), 347-370.
- Sam, M., Li, H.C. (2004), Gold price, neural networks and genetic algorithm. Computational Economics, 23(2), 193-200.

- Sari, R., Hammoudeh, S., Soytas, U. (2010), Dynamics of oil price, precious metal prices, and exchange rate. Energy Economics, 32, 351-362.
- Smith, G. (2001), The price of gold and stock price indices for the United States. Available from: http://www.spdrgoldshares.com/media/GLD/file/Gold&USStockIndicesDEC200120fina.pdf.
- Smith, G. (2002), London gold prices and stock price Indices in Europe and Japan. Available from: http://www.spdrgoldshares.com/media/ GLD/file/GOLD&EUJPStockIndicesFeb2002.pdf.
- Soytas, U., Sari, R., Hammoudeh, S., Hacihasanoglu, E. (2009), World oil prices, precious metal prices and macro economy in Turkey. Energy Policy, 37, 5557-5566.
- Taylor, S.J. (1986), Modeling Financial Time Series. Chichester: Wiley.